



Estimating Epidemic Severity Rates

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Time-varying severity rates in epidemiology

- Severity rates express the probability that a primary event at time t will result in serious secondary event, e.g.
 - Case-fatality rate (CFR)
 - Hospitalization-fatality rate (HFR)
- Time-varying or stationary?
 - Most academic work on estimating severity rates assumes stationarity over time.
 - Severity rates constantly change due to new variants, therapeutics, etc.
 - Epidemiologists at the CDC use time-varying rates to analyze new risks.

Newsletters

The Atlantic

Saved

HEALTH

How Many Americans Are About to Die?

A new analysis shows that the country is on track to pass spring's grimmest record.

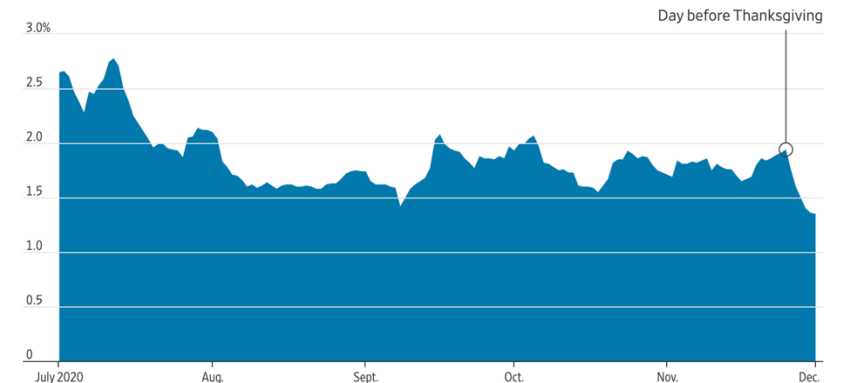
By Alexis C. Madrigal and Whet Moser

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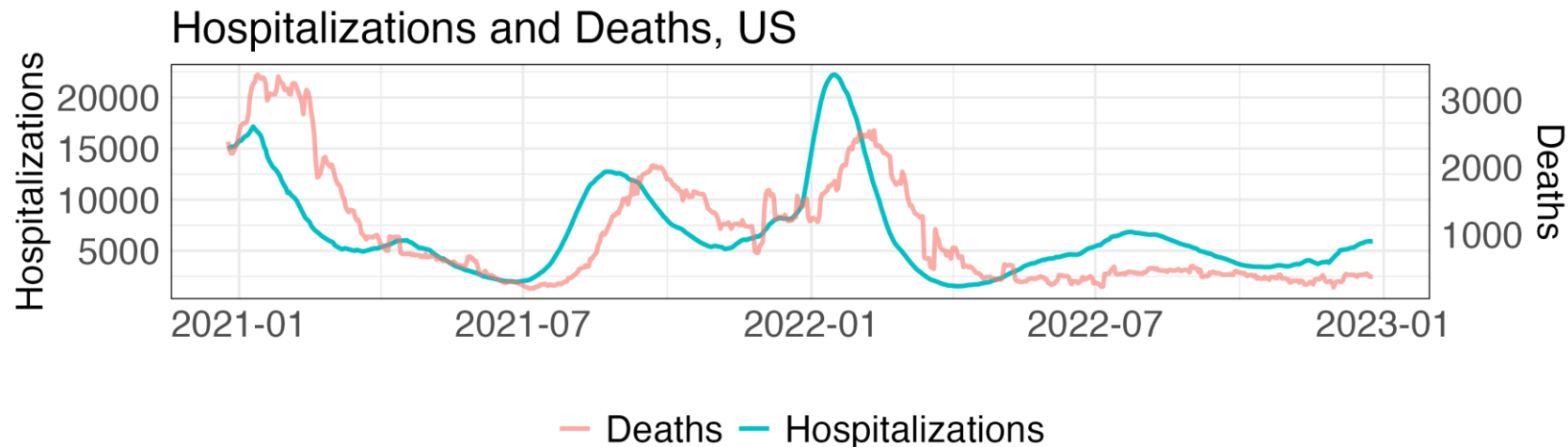
Winter Warning

The U.S. case fatality rate calculated with a 22-day lag between reported cases and deaths points to wave of new fatalities ahead



Often estimate severity from aggregate data

- Calculating severity rates is straightforward with a line list of patient outcomes.
 - CFR: Observe fraction of patients that tested positive at t who ultimately die.
- Maintaining such a line list may be unrealistic or impossible
 - In this case, severity rates must be estimated from aggregate count data.



Standard ratio estimators

- Most estimators for severity rates are simple ratios (“case fatality ratio”) between secondary events and at-risk primary events
- The standard time-varying approach is a lagged ratio of aggregate counts:

$$\widehat{\text{CFR}}_t = \frac{\text{Deaths at } t}{\text{Cases at } t - \ell}$$

- A more principled generalization uses the delay distribution:

$$\widehat{\text{CFR}}_t = \frac{\text{Deaths at } t}{\sum_k \{\text{Cases at } t - k\} \times \hat{\mathbb{P}}(\text{Death is at } k \text{ days})}$$

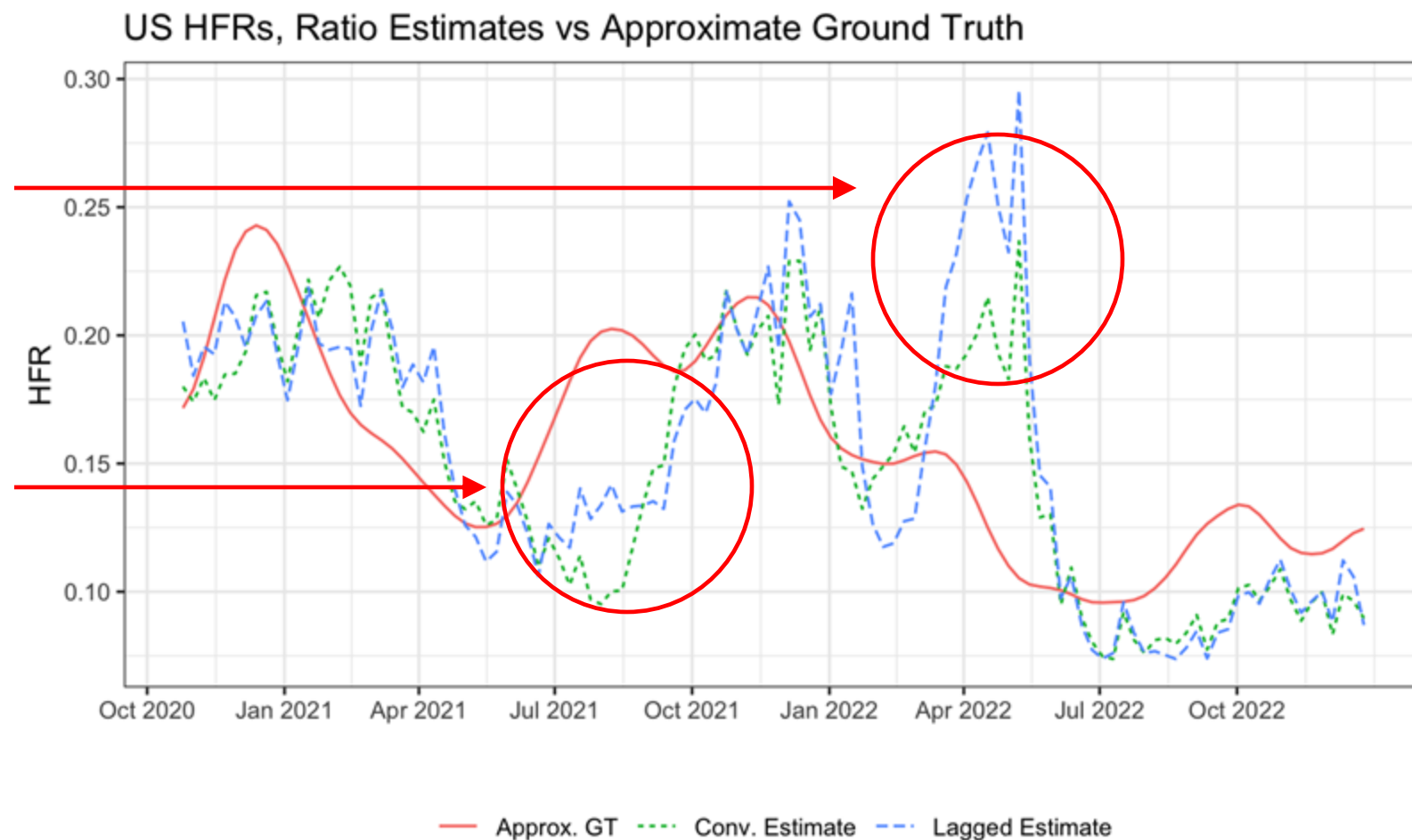
Our work: Understanding the bias of these ratios and proposing statistically sound alternatives.

Observed these ratios exhibit huge bias

Notable failures, HFR:

- Signaled enormous, nonexistent surge after Omicron peak – especially lagged ratio.
- Ignored higher risk as Delta took over

Findings robust across parameters, geography, etc.



Proposed solution: model the relationship between series

- Let $Y_t/X_{s \leq t}$ denote e.g. the number of deaths at time t given prior hospitalizations.

$$Y_t | X_{s \leq t} = \sum_{k=0}^d \sum_{i=1}^{x_{t-k}} \mathbf{1}\{i^{\text{th}} \text{ case at } t-k \text{ died at } t\}$$

- We identify this adheres to a *Poisson Binomial* distribution – a generalization of the binomial distribution where not all success probabilities are equal.
- While its PMF is intractable, it is well-approximated by a Gaussian with mean

$$\mu_t = \sum_{k=0}^d x_{t-k} \mathbb{P}(\text{die at } t \mid \text{hosp at } t-k) = \sum_{k=0}^d x_{t-k} \pi_k p_{t-k}$$

and variance

$$\sigma_t^2 = \sum_{k=0}^d x_{t-k} \pi_k p_{t-k} (1 - \pi_k p_{t-k}) \approx \mu_t.$$

Proposed solution: approximate MLE of probabilistic model

$$\hat{p}_{(t_0-d):T}^{\text{MLE}} = \operatorname{argmax}_p \mathcal{L}(p) = \operatorname{argmin}_p -\log \mathbb{P}(Y_t \forall t | X_{s \leq t} \forall t, \pi, p)$$

Correlation is negligible

$$\approx \operatorname{argmin}_p \sum_{t=t_0}^T -\log \mathbb{P}(Y_t | X_{s \leq t}, \pi, p)$$

Normal approximation at all t

$$\approx \operatorname{argmin}_p \sum_{t=t_0}^T -\log \Phi\left(\frac{Y_t - \mu_t(p)}{\sigma_t(p)}\right)$$

Ignore variance term

$$\approx \operatorname{argmin}_p \sum_{t=t_0}^T \frac{(Y_t - \mu_t(p))^2}{\sigma_t^2(p)}$$

Plug-in variance

$$\approx \operatorname{argmin}_p \sum_{t=t_0}^T \frac{1}{\hat{\mu}_t} \left(Y_t - \sum_{k=0}^d X_{t-k} \pi_k p_{t-k}\right)^2$$

Plug-in delay distribution

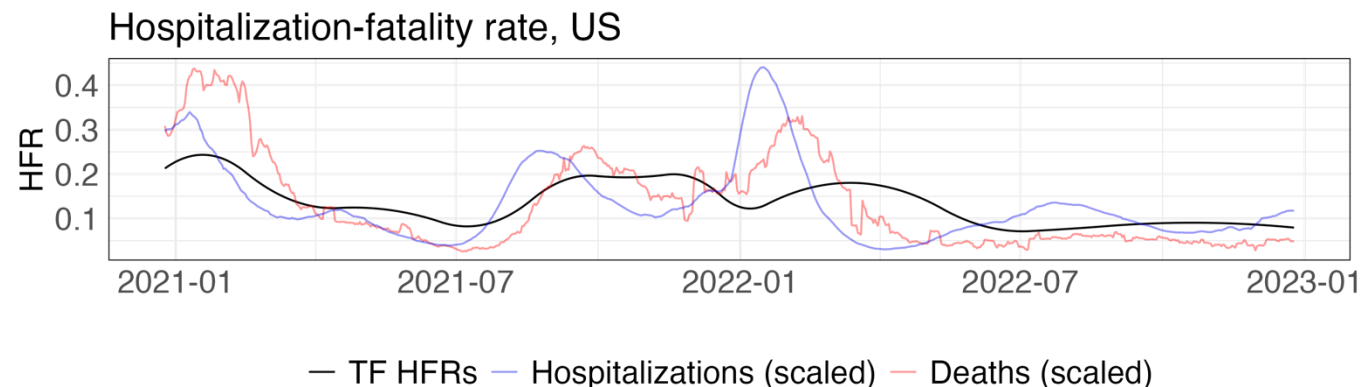
$$\approx \operatorname{argmin}_p \sum_{t=t_0}^T \frac{1}{\hat{\mu}_t} \left(Y_t - \sum_{k=0}^d X_{t-k} \gamma_k p_{t-k}\right)^2.$$

Proposed solution: learn severity rates with smoothed MLE

- To find a smooth solution for this overparameterized problem, we maximize the likelihood subject to a *trend filtering* penalty.

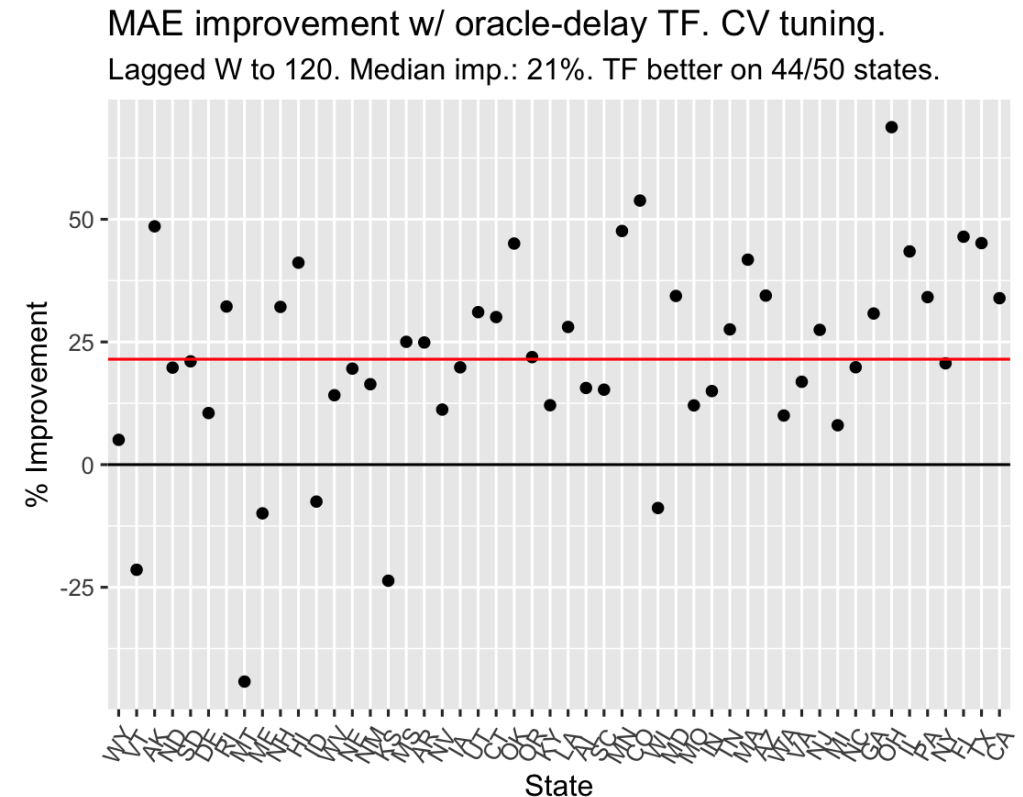
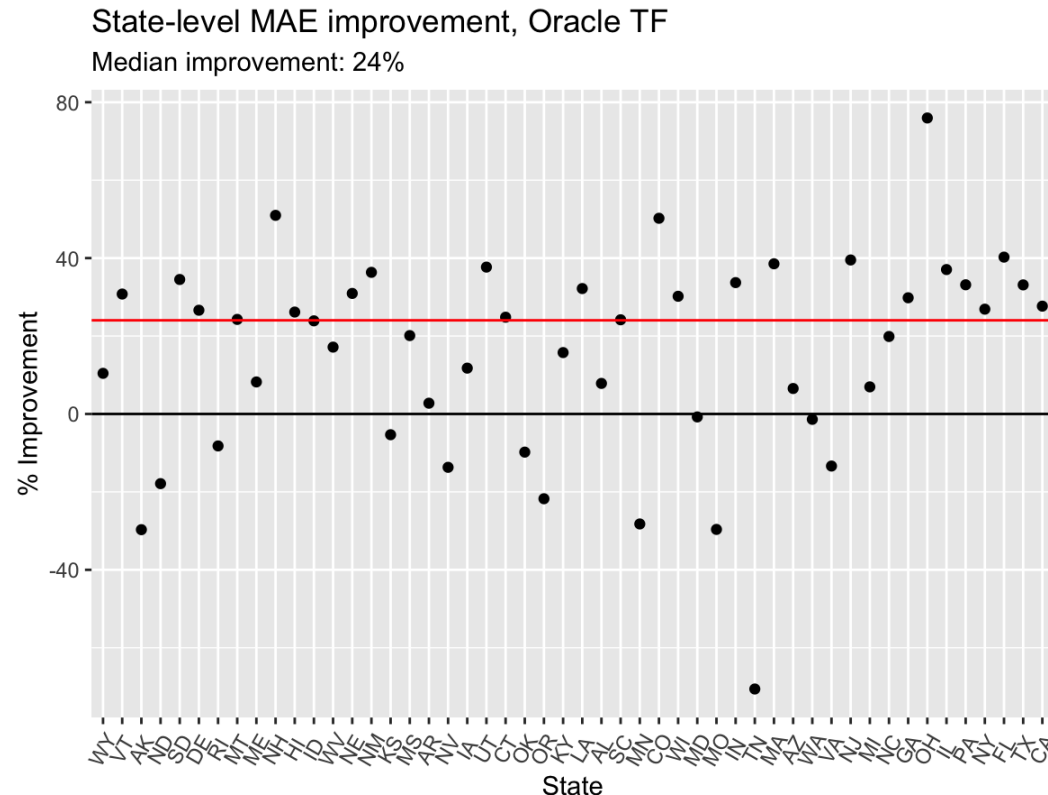
$$\hat{p}^{\text{TF}} = \operatorname{argmin}_{p \geq 0} \sum_{t=t_0}^T \frac{1}{\hat{\mu}_t} (Y_t - \sum_{j=0}^d X_{t-j} \gamma_j p_{t-j})^2 + \lambda \|D^{(k+1)} p\|_1$$

- The difference matrix $D^{(k+1)}$ contains finite differencing operations of order $k+1$. The L1 penalty encourages p to have sparse $k+1^{\text{th}}$ discrete derivatives, so solutions are piecewise polynomials of order k .
 - Trend filtering is more locally adaptive than smoothing splines.



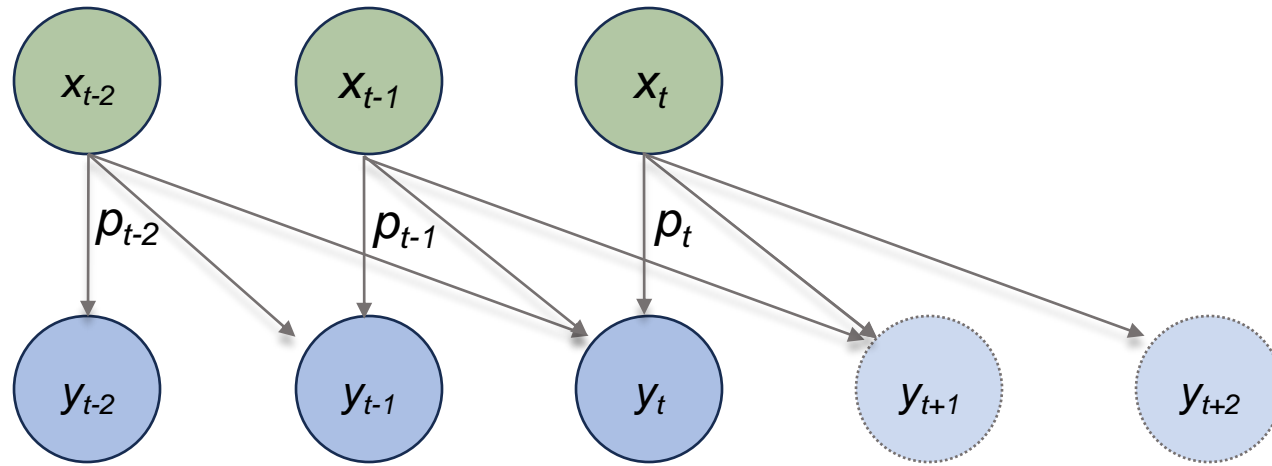
Trend filtering estimator outperforms lagged estimator

- State-level deaths simulated from overdispersed probabilistic model.
- On average, trend filtering **lowers MAE by >20%** over the lagged estimator – with both cross validation and oracle tuning.



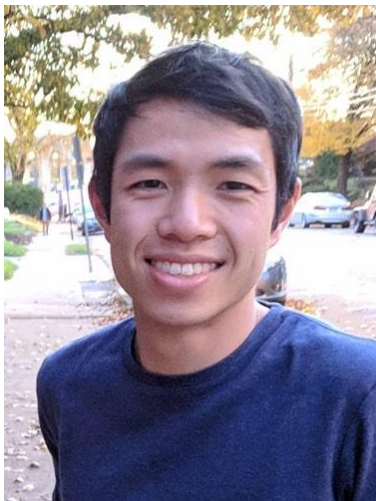
Ongoing: Adapt trend filtering for real-time setting

- Requires extra regularization to mitigate tail variability, since most recent severity rates used for fewer observed predictions.



- Jahja et al. (2022) used natural trend filtering & tapered smoothing for similar deconvolution problem.
- Also aim to quantify uncertainty of severity estimates and compare to convolutional ratio.

Collaborators





*Thanks for
your attention!*